# Area Under the ROC Curve and the Average Precision: A closer look 

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## Outline

- Motivation
- What is average precision
- The hit curve
- The AP and AUC connection
- AP and its variance
- Examples
- Summary and future work


## Motivating Data

Digital Mammography Imaging Screening Trial (Pisano et al. 2005 New England Journal of Medicine)

| Malignancy score |  | 7 | 6 | 5 | 4 | 3 | 2 | 1 | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Digital | Category Total | 11 | 29 | 69 | 1061 | 2224 | 6588 | 32588 | 42570 |
|  | Cancers | 10 | 18 | 25 | 85 | 49 | 25 | 122 | 334 |
| Film | Category <br> Total | 17 | 29 | 70 | 942 | 2291 | 6910 | 32486 | 42745 |
|  | Cancers | 13 | 24 | 25 | 74 | 35 | 33 | 131 | 335 |

## Predicting the Rare Class

- Cancer screening: detect from the asymptomatic population the diseased subjects, who make up a very small proportion (typically < 1\%).
- Drug discovery: identify potential chemical compounds that are biologically active for some target (typically < 5\%).
- Information retrieval


## Performance Measures for Classifiers

- Threshold Dependent Measure
- Misclassification rate
- Sensitivity and Specificity
- Positive and Negative Predictive Value
- Threshold Independent Measure
- Area Under the ROC* Curve (AUC)
- Average precision


## Average Precision

- Definition
$\left\{Y_{(1)}, Y_{(2)}, Y_{(3)}, \ldots Y_{(m)}, \ldots, Y_{(n)}\right\}$. where $Y$ is the class label.
Precision at $Y_{(m)}: c_{m}=\frac{\sum_{i=1}^{m} Y_{(i)}}{m}$ (i.e. the proportion of class 1 subjects in the top $m$ ranked subjects)

$$
\begin{aligned}
& \text { Average Precision }=\frac{1}{n_{1}} \sum_{m=1}^{n} Y_{(m)} c_{m} \\
& \qquad \text { where } n_{1}=\sum_{m=1}^{n} Y_{(m)}
\end{aligned}
$$

## An Illustration Example

| Rank | Classifier 1 | Classifier 2 | Classifier 3 |
| :--- | :--- | :--- | :--- |
| 1 | 1 | 1 | 0 |
| 2 | 1 | 0 | 0 |
| 3 | 1 | 1 | 1 |
| 4 | 0 | 0 | 1 |
| 5 | 0 | 1 | 1 |
| AP | 1 | 0.76 | 0.48 |



Notations

$$
\begin{aligned}
\pi & =P(Y=1) \\
r & =P(X>x)=F_{X}(x) \\
h(r) & =P(X>x, Y=1)=\pi F_{1}(x)
\end{aligned}
$$

ROC curve


$$
\begin{aligned}
\pi & =\mathrm{P}(\mathrm{Y}=1) \\
\mathrm{r} & =\mathrm{P}(\mathrm{X}>\mathrm{x}) \\
\mathrm{h}(\mathrm{r}) & =\mathrm{P}(\mathrm{X}>\mathrm{x}, \mathrm{Y}=1)=\pi \mathrm{F}_{1}(\mathrm{x}) \\
A U C & \stackrel{\text { def }}{=} \int_{0}^{1} \frac{h(r)}{\pi} d\left[\left.\frac{r-h(r)}{1-\pi}\right|^{2}\right. \\
& =\frac{1}{\pi(1-\pi)}\left[\int_{0}^{1} h(r) d r-\frac{\pi^{2}}{2}\right]
\end{aligned}
$$

Hit curve


$$
\begin{aligned}
A P & \stackrel{\text { def }}{=} \int_{0}^{1} \frac{h(r)}{r} d \frac{h(r)}{\pi} \\
& =\frac{1}{\pi} \int_{0}^{1} \frac{h(r)}{r} d h(r)
\end{aligned}
$$

Hit curve


Approximate the hit curve by a quasi-concave curve, let $\beta$ be the initial true positive rate of the underlying test

$$
\mathrm{h}(r)=\left\{\begin{array}{lr}
\beta r, & r \in[0, \alpha] \\
\frac{\pi-\alpha \beta}{1-\alpha}(r-\alpha)+\alpha \beta, & r \in(\alpha, 1]
\end{array}\right.
$$



Theorem 1: If two hit curves, $h_{1}(r)$ and $h_{2}(r)$, both belong to the quasi-concave family, and are parameterized respectively by $\left(\alpha_{1}, \beta_{1}\right)$ and $\left(\alpha_{2}, \beta_{2}\right)$, then $\operatorname{AUC}\left(h_{1}\right)=\operatorname{AUC}\left(h_{2}\right)$ if and only if

$$
\left(\beta_{1}-\pi\right) \alpha_{1}=\left(\beta_{2}-\pi\right) \alpha_{2}
$$

Theorem 2: If a hit curve, $h(r)$, belongs to the quasiconcave family, then

$$
\widetilde{A P}(h) \approx \beta \times \widetilde{A U C}(h)
$$

where AP and AUC are re-scaled to lie between 0 and 1 for any hit curve $h$

$$
\begin{gathered}
\widetilde{\mathrm{AP}} \equiv \frac{\mathrm{AP}-\pi}{1-\pi} \\
\widetilde{\mathrm{AUC}} \equiv \frac{\mathrm{AUC}-1 / 2}{1-1 / 2}=2 \mathrm{AUC}-1 .
\end{gathered}
$$

## Simulation Study

- Non-diseased subjects $(Y=0), f_{0}(x) \sim N(0,1)$
- Diseased subjects $(Y=1), f_{1}(x) \sim N(\Delta, 1)$
- Simulation settings:

$$
\begin{aligned}
& -\Delta=0.5 \text { or } 2 \\
& -\pi=0.1 \text { or } 0.5 \\
& -\mathrm{n}=500
\end{aligned}
$$



## AP vs. AUC in Ordinal Data

- Radiologist reading of an image
- Clinical symptom
- Psychology questionnaire

| Score | $x_{1}$ | $>$ | $x_{2}$ | $>\cdots>$ | $x_{k}$ | $>$ | $x_{k+1}$ | $>\cdots>$ | $x_{K}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Partition | $R_{1}$ | $R_{2}$ | $\cdots$ | $R_{k}$ | $R_{k+1}$ | $\cdots$ | $R_{K}$ | Total |  |  |
| Class-1 | $\bar{Z}_{1}$ | $\bar{Z}_{2}$ | $\cdots$ | $\bar{Z}_{k}$ | $\bar{Z}_{k+1}$ | $\cdots$ | $\bar{Z}_{K}$ | $n_{1}$ |  |  |
| Class-0 | $\bar{Z}_{1}$ | $\bar{Z}_{2}$ | $\cdots$ | $\bar{Z}_{k}$ | $\bar{Z}_{k+1}$ | $\cdots$ | $\bar{Z}_{K}$ | $n_{0}$ |  |  |
| Total | $S_{1}$ | $S_{2}$ | $\cdots$ | $S_{k}$ |  | $S_{k+1}$ | $\cdots$ | $S_{K}$ | $n$ |  |

$$
\begin{aligned}
\text { AP } & =\underbrace{\left[\frac{Z_{1}}{S_{1}}\right]}_{w_{1}}\left[\frac{Z_{1}}{n_{1}}\right]+\underbrace{\left[\frac{Z_{1}+Z_{2}}{S_{1}+S_{2}}\right]}_{w_{2}}\left[\frac{Z_{2}}{n_{1}}\right]+\ldots+\underbrace{\left[\frac{Z_{1}+Z_{2}+\ldots+Z_{K}}{S_{1}+S_{2}+\ldots+S_{K}}\right]}_{w_{K}}\left[\frac{Z_{K}}{n_{1}}\right] \\
& =\sum_{k=1} w_{k}\left[\frac{Z_{k}}{n_{1}}\right] .
\end{aligned}
$$

$$
\begin{aligned}
\mathrm{AUC} & =\frac{n}{n_{0}}\{\underbrace{\left[\frac{S_{1}+S_{2}+\ldots+S_{K}}{n}\right]}_{w_{1}^{\prime}}\left[\frac{Z_{1}}{n_{1}}\right]+\underbrace{\left[\frac{S_{2}+\ldots+S_{K}}{n}\right]}_{w_{2}^{\prime}}\left[\frac{Z_{2}}{n_{1}}\right]+\ldots+\underbrace{\left[\frac{S_{K}}{n}\right]}_{w_{K}^{\prime}}\left[\frac{Z_{K}}{n_{1}}\right]\}-\frac{1}{2}\left(\frac{n_{1}}{n_{0}}\right) \\
& =\frac{n}{n_{0}} \sum_{k=1} w_{k}^{\prime}\left[\frac{Z_{k}}{n_{1}}\right]-\frac{1}{2}\left(\frac{n_{1}}{n_{0}}\right)
\end{aligned}
$$

## A Simulated Example



Weights, $w_{k}$ for AP and $w_{k}^{\prime}$ for AUC, in a simulated example. $f_{0}(x) \sim N(0,1)$ and $f_{1}(x) \sim N(\Delta, 1)$ where $\Delta_{A}=1$ and $\Delta_{B}=0.25 ; \pi=0.1$.

## MLE of AP

| Score | $x_{1}$ | $>$ | $x_{2}$ | $>\cdots>$ | $x_{k}$ | $>$ | $x_{k+1}$ | $>\cdots>$ | $x_{K}$ |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Partition | $R_{1}$ | $R_{2}$ | $\cdots$ | $R_{k}$ | $R_{k+1}$ | $\cdots$ | $R_{K}$ | Total |  |  |
| Class-1 | $Z_{1}$ | $Z_{2}$ | $\cdots$ | $Z_{k}$ | $Z_{k+1}$ | $\cdots$ | $Z_{K}$ | $n_{1}$ |  |  |
| Class-0 | $\bar{Z}_{1}$ | $\bar{Z}_{2}$ | $\cdots$ | $\bar{Z}_{k}$ | $\bar{Z}_{k+1}$ | $\cdots$ | $\bar{Z}_{K}$ | $n_{0}$ |  |  |
| Total | $S_{1}$ | $S_{2}$ | $\cdots$ | $S_{k}$ | $S_{k+1}$ | $\cdots$ | $S_{K}$ | $n$ |  |  |

Data in the $2 \times \mathrm{K}$ table follow

$$
\begin{aligned}
\left(Z_{1}, Z_{2}, \ldots, Z_{K}\right) \mid n_{1} & \sim \operatorname{multinomial}\left(n_{1} ; p_{1}, p_{2}, \ldots, p_{K}\right) \\
\left(\bar{Z}_{1}, \bar{Z}_{2}, \ldots, \bar{Z}_{K}\right) \mid n_{1} & \sim \operatorname{multinomial}\left(n-n_{1} ; q_{1}, q_{2}, \ldots, q_{K}\right) \\
n_{1} & \sim \operatorname{binomial}(n, \pi)
\end{aligned}
$$

where

$$
p_{k}=\int_{R_{k}} f_{1}(x) d x, \quad q_{k}=\int_{R_{k}} f_{0}(x) d x
$$

## Asymptotic Variance of AP

$$
\widehat{\mathrm{AP}}=g\left(\widehat{p_{k}}, \widehat{q_{k}}, \widehat{\pi}\right)=\sum_{k=1}^{K}\left[\widehat{\hat{p}_{k}}\left(\frac{\widehat{\pi} \sum_{k^{\prime} \leq k} \widehat{p}_{k^{\prime}}}{\widehat{\pi} \sum_{k^{\prime} \leq k} \widehat{p}_{k^{\prime}}+(1-\widehat{\pi}) \sum_{k^{\prime} \leq k} \widehat{\hat{q}_{k^{\prime}}}}\right)\right]
$$

Apply the Delta method, we get

$$
\widehat{\operatorname{var}}(\widehat{A P}) \approx(\nabla g)^{T} \hat{\jmath}^{-1}(\nabla g)
$$

## Example 1

Digital Mammography Imaging Screening Trial (Pisano et al. 2005 New England Journal of Medicine)

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42,760 screening participants underwent two screening technology, 335 were diagnosed with breast cancer at 15 months follow-up.

Given that 335 breast cancer diagnosed in 42,760 screening participants at 15 months follow-up, the prevalence $\pi$ is 0.00783 .

## Seven-point Malignancy Scale

## Film mammography <br> 0.735 (0.012) 0.166 (0.022) <br> Digital mammography $\quad 0.753$ ( 0.012 ) 0.144 (0.021)

Remark: Resampling method can be used for the inference of the difference in AP when we have paired data.

## Example 2

Mass spectrometry data for prostate cancer (Adam et al. 2002 Cancer Research)

- 779 potential biomarkers were assessed in 83 late-stage prostate cancer patients and 82 normal subjects.

|  |  | Standard Error of AP |  |  |
| ---: | :---: | :---: | :---: | :---: |
| Biomarker | AP | Asymptotic | P-Bootstrap | NP-Bootstrap |
| 3896.641 | 0.878 | 0.0345 | 0.0344 | 0.0344 |
| 8355.562 | 0.856 | 0.0336 | 0.0339 | 0.0340 |
| 8141.232 | 0.850 | 0.0319 | 0.0324 | 0.0321 |
| 8295.641 | 0.833 | 0.0328 | 0.0327 | 0.0327 |
| 5074.164 | 0.833 | 0.0403 | 0.0405 | 0.0403 |
| 4071.184 | 0.831 | 0.0368 | 0.0364 | 0.0366 |
| 6949.220 | 0.824 | 0.0414 | 0.0415 | 0.0413 |
| 9149.121 | 0.822 | 0.0378 | 0.0380 | 0.0378 |






## A Thought Experiment

- The biomarker study is based on a case-control study with the intention to identify potential screening markers.
- How AP and the ranking of biomarkers is affected when the prevalence is much lower as in a screening setting?
Inflate the controls by replicating them

| Biomarker | AUC | AP |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $n_{0} \times 1$ | $n_{0} \times 1$ | $n_{0} \times 10$ | $n_{0} \times 100$ |
| $\mathbf{8 3 5 5 . 5 6 2}$ | $\mathbf{0 . 8 4 9}$ | 0.856 | 0.606 | 0.571 |
| $\mathbf{7 8 1 9 . 7 5 1}$ | $\mathbf{0 . 8 5 0}$ | 0.802 | 0.370 | 0.062 |
| 6949.220 | 0.881 | $\mathbf{0 . 8 2 4}$ | 0.452 | 0.205 |
| $\mathbf{9 1 4 9 . 1 2 1}$ | 0.832 | $\mathbf{0 . 8 2 2}$ | 0.512 | 0.225 |

## Summary

- A single numerical measure, independent of threshold
- Connection between AP and AUC
- Empirical estimation of AP and its asymptotic variance
- Practical relevance


## Future work

- Asymptotic variance of AP when test score is continuous
- Assessing risk predictor
- Assessing survival models

