

STAT 312 Lab 11

1. Define a function $f : \mathbb{R}^n \rightarrow \mathbb{R}^1$ by $f(\mathbf{x}) = \mathbf{x}'\mathbf{A}\mathbf{x}$, where \mathbf{A} is a symmetric matrix. Calculate:

- (a) the Jacobian $J_f(\mathbf{x})$,
- (b) the gradient $\nabla_f(\mathbf{x})$,
- (c) the Hessian $H_f(\mathbf{x})$.

2. Consider the problem of estimating a parameter vector β , by Least Squares, in the linear model $\mathbf{y} = \mathbf{X}\beta + \text{random error}$. Here $\mathbf{y} : n \times 1$ and $\mathbf{X} : n \times p$ are constants and \mathbf{X} has rank $p < n$. Suppose that the parameters are required to satisfy q independent linear constraints of the form $\mathbf{A}\beta = \mathbf{0}_{q \times 1}$, where $\mathbf{A}_{q \times p}$ has rank q . Thus the mathematical problem is

$$\text{Minimize } S(\beta) = \|\mathbf{y} - \mathbf{X}\beta\|^2 \text{ over } \beta \in R^p, \text{ subject to constraints } \mathbf{A}\beta = \mathbf{0}_{q \times 1}.$$

Show that the solution to this problem is

$$\hat{\beta} = \left[\mathbf{I} - (\mathbf{X}'\mathbf{X})^{-1}\mathbf{A}'(\mathbf{A}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{A}')^{-1}\mathbf{A} \right] \hat{\beta}_{LS},$$

where $\hat{\beta}_{LS} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$ is the unconstrained LS estimator.

3. The following problem is of interest in Mathematical Finance. A portfolio consists of n investments, whose possible returns can be viewed as random variables. The i^{th} investment has a mean return of μ_i , a variance of σ_{ii} , and the covariance between investments i and j is σ_{ij} . Define $\boldsymbol{\mu}_{n \times 1} = (\mu_1, \dots, \mu_n)'$ and $\boldsymbol{\Sigma}_{n \times n} = (\sigma_{ij})$. Suppose that one invests a fraction w_i of the total (which we take as = 1, in appropriate units) into investment i . Define $\mathbf{w}_{n \times 1} = (w_1, \dots, w_n)'$.

- (a) Show that the expected return on this portfolio is $\boldsymbol{\mu}'\mathbf{w}$, and that the variance is $\mathbf{w}'\boldsymbol{\Sigma}\mathbf{w}$.
- (b) Suppose that one requires an expected return of μ_0 . Derive the weights which will accomplish this, with a minimum possible variance. [*Hint*: this is a version of Problem (P) in Lecture 31.]
- (c) Explain why (b) is only a reasonable question if μ_0 lies between the smallest and largest of the μ_i .

...over

4. Recall from Lab 6 that the density of $X = Z^2$, where $Z \sim N(0, 1)$, is $f_1(x) = \frac{(\frac{x}{2})^{\frac{1}{2}-1}}{2\Gamma(\frac{1}{2})}e^{-x/2}$, the χ_1^2 density. Show that the χ_n^2 density is $f_n(x) = \frac{(\frac{x}{2})^{\frac{n}{2}-1}}{2\Gamma(\frac{n}{2})}e^{-x/2}$. [*Hint*: What r.v. has this density? Use m.g.f.s.]
5. Suppose that X_1, \dots, X_n are i.i.d. $N(\mu, \sigma^2)$ r.v.s.
- Obtain the mles of μ and σ^2 .
 - Derive the information matrix.
 - What is the joint asymptotic distribution of the two mles? [*Note*: Expression (34.2) in the notes is usually the easier of the two ways to calculate the Information matrix.]
 - Obtain the mle of σ^2 if μ is known.
 - Is either mle of σ^2 unbiased?
6. Let X_1, \dots, X_n be a sample from the Uniform density

$$\begin{aligned} p(x; \theta) &= \begin{cases} \frac{1}{\theta}, & 0 \leq x \leq \theta, \\ 0, & \text{otherwise} \end{cases} \\ &= \frac{1}{\theta} I(0 \leq x \leq \theta). \end{aligned}$$

Determine the MLE of θ . [*Hint*: This is a problem in which the ‘regularity conditions’ – being able to differentiate the likelihood, etc. – don’t hold. Just plot the likelihood function to see where it is maximized.]