

## STAT 312 Lab 11

1. Define a function  $f : \mathbb{R}^n \rightarrow \mathbb{R}^1$  by  $f(\mathbf{x}) = \mathbf{x}'\mathbf{A}\mathbf{x}$ , where  $\mathbf{A}$  is a symmetric matrix. Calculate:
  - (a) the Jacobian  $J_f(\mathbf{x})$ ,
  - (b) the gradient  $\nabla_f(\mathbf{x})$ ,
  - (c) the Hessian  $H_f(\mathbf{x})$ .

2. Consider the problem of estimating a parameter vector  $\beta$ , by Least Squares, in the linear model  $\mathbf{y} = \mathbf{X}\beta + \text{random error}$ . Here  $\mathbf{y} : n \times 1$  and  $\mathbf{X} : n \times p$  are constants and  $\mathbf{X}$  has rank  $p < n$ . Suppose that the parameters are required to satisfy  $q$  independent linear constraints of the form  $\mathbf{A}\beta = \mathbf{0}_{q \times 1}$ , where  $\mathbf{A}_{q \times p}$  has rank  $q$ . Thus the mathematical problem is

*Minimize*  $S(\beta) = \|\mathbf{y} - \mathbf{X}\beta\|^2$  *over*  $\beta \in R^p$ , *subject to constraints*  $\mathbf{A}\beta = \mathbf{0}_{q \times 1}$ .

Show that the solution to this problem is

$$\hat{\beta} = \left[ \mathbf{I} - (\mathbf{X}'\mathbf{X})^{-1}\mathbf{A}'(\mathbf{A}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{A}')^{-1}\mathbf{A} \right] \hat{\beta}_{LS},$$

where  $\hat{\beta}_{LS} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$  is the unconstrained LS estimator.

3. The following problem is of interest in Mathematical Finance. A portfolio consists of  $n$  investments, whose possible returns can be viewed as random variables. The  $i^{\text{th}}$  investment has a mean return of  $\mu_i$ , a variance of  $\sigma_{ii}$ , and the covariance between investments  $i$  and  $j$  is  $\sigma_{ij}$ . Define  $\boldsymbol{\mu}_{n \times 1} = (\mu_1, \dots, \mu_n)'$  and  $\boldsymbol{\Sigma}_{n \times n} = (\sigma_{ij})$ . Suppose that one invests a fraction  $w_i$  of the total (which we take as = 1, in appropriate units) into investment  $i$ . Define  $\mathbf{w}_{n \times 1} = (w_1, \dots, w_n)'$ .
  - (a) Show that the expected return on this portfolio is  $\boldsymbol{\mu}'\mathbf{w}$ , and that the variance is  $\mathbf{w}'\boldsymbol{\Sigma}\mathbf{w}$ .
  - (b) Suppose that one requires an expected return of  $\mu_0$ . Derive the weights which will accomplish this, with a minimum possible variance. [*Hint*: this is a version of Problem (P) in Lecture 31.]
  - (c) Explain why (b) is only a reasonable question if  $\mu_0$  lies between the smallest and largest of the  $\mu_i$ .

...over

4. Recall from Lab 6 that the density of  $X = Z^2$ , where  $Z \sim N(0, 1)$ , is  $f_1(x) = \frac{(\frac{x}{2})^{\frac{1}{2}-1}}{2\Gamma(\frac{1}{2})}e^{-x/2}$ , the  $\chi_1^2$  density. Show that the  $\chi_n^2$  density is  $f_n(x) = \frac{(\frac{x}{2})^{\frac{n}{2}-1}}{2\Gamma(\frac{n}{2})}e^{-x/2}$ . [*Hint*: What r.v. has this density? Use m.g.f.s.]
5. Suppose that  $X_1, \dots, X_n$  are i.i.d.  $N(\mu, \sigma^2)$  r.v.s.
- Obtain the mles of  $\mu$  and  $\sigma^2$ .
  - Derive the information matrix.
  - What is the joint asymptotic distribution of the two mles? [*Note*: Expression (34.2) in the notes is usually the easier of the two ways to calculate the Information matrix.]
  - Obtain the mle of  $\sigma^2$  if  $\mu$  is known.
  - Is either mle of  $\sigma^2$  unbiased?
6. Let  $X_1, \dots, X_n$  be a sample from the Uniform density

$$\begin{aligned} p(x; \theta) &= \begin{cases} \frac{1}{\theta}, & 0 \leq x \leq \theta, \\ 0, & \text{otherwise} \end{cases} \\ &= \frac{1}{\theta} I(0 \leq x \leq \theta). \end{aligned}$$

Determine the MLE of  $\theta$ . [*Hint*: This is a problem in which the ‘regularity conditions’ – being able to differentiate the likelihood, etc. – don’t hold. Just plot the likelihood function to see where it is maximized.]